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THEMIS SIGNAL ANALYSIS STATISTICS RESEARCH PROGRAM

ASYMPTOTIC INDEPENDENCE BETWEEN LARGEST AND SMALLEST OF A SET OF INDEPENDENT OBSERVATIONS

by

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Technical Report No. 17
Department of Statistics THEMIS Contract

Department of Statistics
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Dallas, Texas 75222

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October 4, 1968

Research sponsored by the Office of Naval Research
Contract N00014-68-A-0515
Project NR 042-260

DEPARTMENT OF STATISTICS
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ABSTRACT

Let X_n and X_1 be the largest and smallest order statistics, respectively, of a set of n independent univariate observations. Under rather general conditions, with respect to the distributions of the individual observations, X_n and X_1 are asymptotically independent. That is, the maximum difference between $P(X_1 \le x_1, X_n \le x_n)$ and $P(X_1 \le x_1)P(X_n \le x_n)$ tends to zero as $n \to \infty$. However, asymptotic independence does not occur for all cases.

INTRODUCTION AND RESULTS

Asymptotic independence of the largest and smallest order statistics of a random univariate sample is well known. The question arises as to what extent this asymptotic independence remains when the observations are still required to be independent but can have arbitrarily different distributions. That is, let X_n and X_1 be the largest and smallest, respectively of a set of n independent observations. When does the maximum of

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* Research partially supported by NASA Grant NGR 44-007-028. Also associated with ONR Contract N00014-68-A-0515.

$$P(X_1 \le x_1)P(X_n \le x_n) - P(X_1 \le x_1, X_n \le x_n), \tag{1}$$

over x_1 and x_n, tend to zero as $n \to \infty$?

The interest is in the range of x_1 values that are meaningful for $P(X_1 \le x_1)$ and range of x_n values that are meaningful for $P(X_n \le x_n)$. That is, the analysis is made in terms of (attainable) percentiles for $P(X_1 \le x_1)$ and for $P(X_n \le x_n)$. Let

$$P(X_n \le x_n) = e^{-a}, \quad P(X_1 \le x_1) = 1 - e^{-b},$$

where <u>a</u> and <u>b</u> are arbitrary but fixed. Then,

$$\prod_{i=1}^{n} F_i(x_n) = e^{-a}, \quad \prod_{i=1}^{n} [1 - F_i(x_1)] = e^{-b},$$

where $F_i(x)$ is the cumulative distribution function (cdf) for the i-th observation ($i = 1, \ldots, n$).

Now consider the $F_i(x_n)$ and $F_i(x_1)$. Let these cdf's be expressed as

$$F_i(x_n) = e^{-a_i/n}, \quad F_i(x_1) = 1 - e^{-b_i/n}$$

where $a_i = a_i(n)$ and $b_i = b_i(n)$. Asymptotic independence between X_n and X_1 always occurs if

$$a_i \le A(n), \qquad b_i \le B(n)$$

for all i. Here, $A(n)$ and $B(n)$ are $O(n)$ and at least one of them is $o(n)$. That is both $A(n)/n$ and $B(n)/n$ tend to constants as $n \to \infty$, and at least one of these constants is zero. For example, the forms $C_1 n/\log n$ and $C_2 n^{1-\epsilon}$ (with $\epsilon > 0$ and fixed but as small as desired) yield zero constants.

Examination shows that asymptotic independence fails to occur only when, for one or more values of i, both a_i and b_i are $O(n)$ and neither is $o(n)$. For these observations, the values of $1 - F_i(x_n)$ and $F_i(x_1)$, representing the "tail" probabilities, are relatively much larger than these values for the other observations (ratio becomes infinite as $n \to \infty$). Thus, for large n, approximate independence of X_n and X_1 should not be accepted when a few of the distributions seem to have a much wider spread (in both tails) than the others.

The next and final section contains derivations of the results that are stated in this section.

DERIVATIONS

The difference (1) can be written

$$\prod_{i=1}^{n} F_i(x_n)[1 - F_i(x_1)] - \prod_{i=1}^{n}[F_i(x_n) - F_i(x_1)]$$

$$= e^{-(a+b)} - e^{-a} \prod_{i=1}^{n}\left[1 - e^{a_i/n} + e^{-(b_i - a_i)/n}\right]$$

$$= e^{-(a+b)} - e^{-a} \prod_{i=1}^{n}\left[1 - \frac{b_i}{n} - \sum_{k=2}^{\infty} \frac{a_i^{\,k} + (-1)^{k+1}(b_i - a_i)^k}{k!\; n^k}\right]$$

$$= e^{-(a+b)} - e^{-a} \exp\left\{\sum_{i=1}^{n} \log_e \left[1 - \frac{b_i}{n} - \sum_{k=2}^{\infty} \frac{a_i^{\,k} + (-1)^{k+1}(b_i - a_i)^k}{k!\; n^k}\right]\right\}$$

$$= e^{-(a+b)}\left(1 - \exp\left\{-\sum_{i=1}^{n} \frac{a_i b_i}{n^2}\left[1 + \frac{a_i + b_i}{2n} + \sum_{k=4}^{\infty} G_{k-2}\left(\frac{a_i}{n}, \frac{b_i}{n}\right)\right]\right\}\right),$$

where G_{k-2} $(a_i/n, b_i/n)$ is a mixed polynomial of degree k - 2 that is symmetrical in a_i/n and b_i/n. The value of

$$\sum_{i=1}^{n} \frac{a_i b_i}{n^2} \left[1 + \frac{a_i + b_i}{2n} + \sum_{k=4}^{\infty} G_{k-2}\left(a_i/n, b_i/n\right) \right] \qquad (2)$$

is largest when some of the a_i have their maximum value, some of the b_i have their maximum value, and the others are zero. Also, the i values are such that the summation of $a_i b_i/n^2$ is largest and, say, A(n) is O(n) with nonzero constant (equivalent results would be obtained if B(n) had the nonzero constant).

Let the values of i such that $a_i = A(n)$ be $i = 1, \ldots, r_n$, while the values such that $b_i = B(n)$ are $i = 1, \ldots, r_1$. Then, since,

$$\sum_{i=1}^{n} a_i = na, \qquad \sum_{i=1}^{n} b_i = nb,$$

r_n is O(1) and r_1 is $O[n/B(n)]$. With these substitutions, (2) becomes

$$r_n[A(n)/n][B(n)/n][1 + O(1)]$$

and tends to zero as $n \to \infty$. Thus, the exponential of the negative of (2) tends to unity and the difference (1) tends to zero.

If both A(n) and B(n) had nonzero constants, the value of (2) would be

$$\min(r_1, r_n)[A(n)/n][B(n)/n][1 + O(1)]$$

and would not tend to zero as $n \to \infty$. Hence the difference (1) would not tend to zero and asymptotic independence does not occur.

UNCLASSIFIED

Security Classification

DOCUMENT CONTROL DATA - R & D

(Security classification of title, body of abstract and indexing annotation must be entered when the overall report is classified)

1. ORIGINATING ACTIVITY (Corporate author) SOUTHERN METHODIST UNIVERSITY	2a. REPORT SECURITY CLASSIFICATION UNCLASSIFIED
	2b. GROUP UNCLASSIFIED

3. REPORT TITLE

Asymptotic Independence Between Largest and Smallest of a Set of Independent Observations

4. DESCRIPTIVE NOTES (Type of report and inclusive dates)

Technical Report

5. AUTHOR(S) (First name, middle initial, last name)

John E. Walsh

6. REPORT DATE October 4, 1968	7a. TOTAL NO. OF PAGES 4	7b. NO. OF REFS 0
8a. CONTRACT OR GRANT NO. N00014-68-A-0515	9a. ORIGINATOR'S REPORT NUMBER(S) 17.	
b. PROJECT NO. NR 042-260		
c.	9b. OTHER REPORT NO(S) (Any other numbers that may be assigned this report)	
d.		

10. DISTRIBUTION STATEMENT

No limitations

11. SUPPLEMENTARY NOTES	12. SPONSORING MILITARY ACTIVITY Office of Naval Research

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DD FORM 1 NOV 65 1473

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